Measurement Invariance

Are we measuring the same thing in different groups?



Today's goal:

If we run a CFA in multiple groups, find out whether we are measuring the same thing in each group

Outline:

- Intro to measurement invariance
- Partial invariance
- Practical example in Mplus (from a paper that is currently under review!)

Types of invariance

Testing different levels of equivalence between groups



Let's say we run a CFA in multiple groups...

- Are measuring the same thing in each group?
- This is called measurement invariance

Why do we want to know this?

Out of fairness

e.g are some IQ-test questions biased against women?



To see if comparisons are warranted

e.g. can we compare measures of satisfaction between cultures?

To detect conceptual differences between groups e.g. do privacy practices have similar meanings in different cultures?

To detect conceptual drift over time

e.g. does "information overload" mean the same thing now as it did 20 years ago?

Types of invariance

Configural invariance (equal form invariance) Does the same model work for both (all) groups?

Metric invariance (equal factor loadings) Are the loadings similar* between groups?

Scalar invariance (equal intercepts/thresholds)

Are the intercepts/thresholds similar between groups? (in most cases this is added on top of metric invariance)

Types of invariance

Equivalence of construct variances/covariances Are the factor variances and correlations the same? (only makes sense on top of metric invariance)

Equivalence of residual variances/covariances

Are the uniquenesses and residual correlations the same?

(also only makes sense on top of metric invariance)

Full equivalence

Can we combine the two groups and run a single model?



Fit two models simultaneously

Allow estimates (loadings, uniqueness, etc.) to be different Do these models both fit reasonably well?





Make the **loadings** equal between groups Does this fit significantly worse than the configural model? If so, metric noninvariance: different contribution per item





Make the **intercepts/thresholds** equal between groups Does this fit significantly worse than the configural model? If so, scalar noninvariance: item is biased





Scalar+Metric

Make **loadings and intercepts/thresholds** equal More common (why?) Usually tested against the metric model





Make **construct variances and covariances** equal Usually done on top of (and tested against) metric model Do the groups equally (co-)vary on the construct?





Make **residual variances (and covariances)** equal Usually done on top of (and tested against) metric model Are the uniquenesses equal across groups?





Make everything equal between groups Essentially, fit a single model on both groups Test vs. more relaxed model (e.g. metric+scalar+construct)





Tests between different levels of equivalence are conducted as chi-square model comparisons

With large N, likely significant!

Solution: also look at approximate fit statistics, especially CFI If CFI differs < .002, essentially no effect If CFI differs < .01, likely no effect



Partial invariance

What if some (but not all) parameters are different?



Question: What if we don't have metric invariance?

- i.e. the metric model is significantly worse than the configural model
- Answer: Go for partial metric invariance!
 - Inspect the loadings in the configural model Which loadings differ the most? These are "differential functioning indicators"



Run the metric model with one parameter relaxed Test against configural model Still significant? Repeat until no longer significant!





You can also start from the configural model Constrain one parameter and test against configural If not significant, continue until it is (then one step back)





Notes:

- You can also do this for partial scalar invariance, and even partial equivalence of construct or residual variances and covariances
- You can combine, and have a model that is e.g. fully metric invariant + partially scalar invariant



The constraints you relax may have theoretical value:

Conceptual differences

- If certain items load differentially, this tells us how the concept differs between groups
- e.g. surveillance-related privacy items load stronger on privacy concern for women than for men



The constraints you relax may have theoretical value:

Biased items

- If certain items have higher/lower intercepts/thresholds, this shows us how the scale may be biased
- e.g. IQ questions that use football metaphors may have a higher intercept for US participants, so such items should be avoided



The constraints you relax may have theoretical value:

Differences in discriminant validity

- If certain factors have much higher correlation in one group, this shows a difference in the complexity of users' attitudes/perceptions/behaviors
- e.g. preventative and collaborative privacy management strategies have a much higher correlation in the US than in Korea and Singapore (to the point where there is a lack of discriminant validity)



Practical example in Mplus



500_dataset.csv: survey of collective privacy management strategies in three countries (the US, Singapore, and Korea)

Columns:

- **–** cntry: 1=US, 2=SG, 3=KR
- male: gender (0=female, 1=male)
- age



Columns (continued):

- csuntag-csauntag: corrective strategies (7 items)
- icconaud-iclimsha: information control strategies (3 items)
- psrespos-pstimrvw: preventive strategies (4 items)
- clnegrul-cledupri: collaborative strategies (7 items)
- nminvite-nmcontac: network control strategies (3 items)
- psuprof-psuaddtg: audience control strategies (6 items)

All 7-point scales; pre-trimmed (all items fit)



invariance.inp

Grouping parameter to indicate the 3 groups

R will automatically run a configural model as specified under "model"

We added some residual covariances, and specified them configurally (separate parameters) for each country

You can do this under model US, model SG, model KR

Note: items not specified as ordered categorical!



invariance.inp

- ANALYSIS: model = configural metric scalar
 - This automatically tests these three types of invariance!
 - (the latter is metric+scalar)



- **− X**²(1156) = 2099.409, p < .0001
- CFI = .935
- **-** ⊤LI = .927
- RMSEA = .070 [.065, .075], p(RMSEA<=.05) < .001



- **X**²(1204) = 2192.533
- Against configural: $\chi^2(48) = 93.124$, p = .0001
- CFI = .932 (decreases .003)
- **-** ⊤LI = .927
- RMSEA = .070 [.066, .075], p(RMSEA<=.05) < .001

No full metric invariance!

But pretty close



- **χ**²(1252) = 2319.573
- Against metric: **X**²(48) = 127.040, p < .0001
- CFI = .927 (decreases .006)
- **-** ⊤LI = .924
- RMSEA = .072 [.067, .076], p(RMSEA<=.05) < .001



metric.inp

The metric model, with modification indices

Result: relax constraint on iclimsha

Determined after some trial and error



partial metric 1.inp

MODEL SG: infoctrl BY iclimsha*; MODEL KR: infoctrl BY iclimsha*;

The metric model, with relaxed constraint on iclimsha

- **− X**²(1202) = 2169.304
- Against configural: $\chi^2(46) = 69.895$, p = .0131
- Modification indices suggest relaxin constraints on icadjcon



partial metric 2.inp

MODEL SG: infoctrl BY iclimsha* icadjcon*; MODEL KR: infoctrl BY iclimsha* icadjcon*;

The metric model, with relaxed constraints on iclimsha and icadjcon

- **χ**²(1200) = 2156.372
- Against configural: $\chi^2(44) = 56.963$, p = .09
- CFI: .934, TLI: .929 (better!), RMSEA: .069 [.065, .074]



iclimsha: I limit what I share on Facebook to only what is appropriate for all of my friends to see Loading US: 0.921, SG: 1.416, KR: 0.650

icadjcon: I adjust the content of my post based on who I think will see it

Loading US: 1.048, SG: 1.523, KR: 1.008

icconaud: Before posting on Facebook I consider the audience that will read my post

Loading US: 1.000, SG: 1.000, KR: 1.000



partial metric full scalar.inp

- Change model to scalar;
- MODEL SG: infoctrl BY iclimsha* icadjcon*; [icadjcon]; [iclimsha];
- MODEL KR: infoctrl BY iclimsha* icadjcon*; [icadjcon]; [iclimsha];
- This frees the intercepts for the freed loadings Why are we doing that?



- **Χ**²(1244) = 2235.565
- Against partial metric: χ²(44) = 79.193, p = .0009
- CFI = .932 (decreases .002)
- **-** ⊤LI = .929
- RMSEA = .069 [.065, .074]
- Modification indices suggest relaxing intercepts of cldispri and cledupri



partial metric partial scalar 2.inp

Add [cldispri] and [cledupri] to MODEL SG and KR

Results:

- **χ**²(1240) = 2212.000
- Against partial metric: **χ**²(40) = 55.628, p = .05
- CFI: .933, TLI: .930 (better!), RMSEA: .069 [.064, .073]



cldispri: Prior to disclosing content, my friends and I discuss the appropriate privacy settings Intercept US: 3.620, SG: 3.740, KR: 3.950

cledupri: I educate my friends about privacy issues Intercept US: 3.849, SG: 3.926, KR: 3.621



cors.inp

Add a test for the correlation between preventive and collaborative strategies

model US: preventive with colstgy (p1); model SG: preventive with colstgy (p2); model KR: preventive with colstgy (p3);

Also get the standardized output



- Wald test: $\chi^2(2)$ = 15.412, p = .0005
- Correlation in US: 0.804 (\sqrt{AVE} of preventive is 0.726!)
- Correlation in SG: 0.584
- Correlation in KR: 0.527

Meaning: US participants do not distinguish between collaborative and preventive strategies!



What if I have a lot of groups (with low N in each group?) Multiple group modeling becomes impossible!

Solution: use a random effect!

- Treat your data like repeated measures
- Create **random** slopes and intercepts for each indicator
- Test whether the variance of these effects is significantly larger than zero

This is called the **alignment** method

"It is the mark of a truly intelligent person to be moved by statistics."

George Bernard Shaw